SOCIAL INTERACTIONS OF MIGRANTS AND TRADE OUTCOMES

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Resumo

Esse artigo analisa o impacto das interações sociais desenvolvidas por imigrantes no comércio bilateral das regiões francesas. Utiliza-se dados da distribuição dos imigrantes nas 95 microrregiões francesas ("département" em francês) em cinco anos para a implementação de uma medida direta do impacto das interações sociais na economia. Comparações diretas confirmam que essa medida é mais completa e robusta do que as medidas usadas na literatura de comércio internacional.

Palavras-chave: Interação social; Comércio internacional; Migração.

Abstract

This paper investigates the impact of social interactions with immigrants on bilateral trade in French regions. Applying data on the distribution of immigrants in 95 French counties ("départment" in French) during 5 years, a measure of formation of networks is proposed and its impact is tested on international trade, providing a direct measure of the impact of social interactions on the economy. Direct comparisons confirm this measure as more complete and robust than measures used in international trade literature.

Keywords: Social Interaction; International Trade; Migration.

JEL classification: F10, F22.

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1 Introduction

Social networks provide a major explanation for the impact immigrants have on international trade.¹ This mechanism relies on two fundamental concepts. Firstly, immigrant communities have the potential to deter violations of informal contracts. Secondly, immigrants benefit from privileged information on home-country and host-country markets.

These interactions are better understood if one considers that newly arrived immigrants benefit from a settled compatriot's community, since they face barriers of language, habits and culture. Work, loans, associates and relationships can be found in such community. Specifically, business opportunities can rise in this context, since market information from both home and host countries is available within the community.

Rauch (1999) points out to the importance for proximity on network contacts. Herander & Saavedra (2005) verify if proximity matters for interactions between community's members residing in the same host country. By analyzing exports in American states, they found that both local immigrants (those living in the exporting state) and nonlocal immigrants (those living in other states) have an impact on exports. Nonetheless, local populations of immigrants have stronger impact on state exports than do nonlocal immigrants, indicating that network effects decrease with distance.

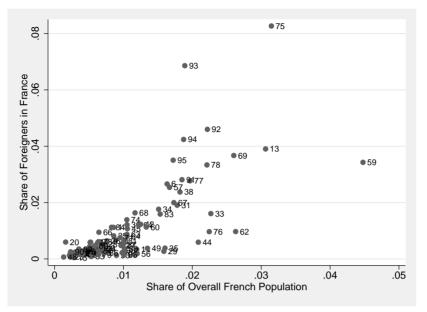
The nature of those interactions, depending in the proximity, is observed in the geographical concentration of foreigners, compared to natives (Bartel 1989). Such interdependence factors are reflected in the distribution of people. Immigrants have a tendency to cluster at higher densities relative to local population. Figure 1 shows the overconcentration of immigrants in France, in comparison with the French population as a whole: 70% of the foreign population² more concentrated than the French population. While there may be specific reasons for that locally, in the region, they do not provide a complete explanation for the overconcentration of immigrants.

This paper investigates the effect between social interactions of immigrants on international trade. Considering the international trade in a certain French region (with a given country), a network's functional form is proposed by counting all immigrants (from this given country) living in France, but taking into account that these effects are a decreasing function of the distance. For instance, it is assumed that a Portuguese in Lyon can contribute to the trade between Portugal and Paris, even if to a lesser extent than could a Portuguese in Paris.

This paper is structured as follows. Section 2 describes the dataset and the variables used. Section 3 reports on the findings before some tentative conclusions are drawn in section 4.

 $^{^1} See$ Rauch (2001) and Wagner et al. (2002) for a synthesis of literature on the impact of immigration on trade

²The foreigner population in France have a coefficient of geographical concentration G=0.24. This positive concentration coefficient equals 0 if foreigners are distributed exactly the same as the total population. Chiswick et al. (2002): "Where the overseas born group has a distribution across regions the same as the total population, G equal to 0. Where the overseas born group is completely segregated, the upper bound of G will equal (100-GS), where GS is the percent of the total population accounted for by the specific birthplace group."



This graph compares the total population of foreigners in France to the total French population. Data is from the 1999 census. Eg: 8% of all the foreigners living in France resided in Paris (coded 75) compared to only 3% of the French population, in this year

Figure 1: Foreigners share versus French share, by French "département"

2 Data and Variables Conception

The migration data used herein derives from the 1968, 1975, 1982, 1990 and 1999 French census. These provide information on the stock of immigrants living in each of the 95 French metropolitan "départements" in those particular years. Total French populations are obtained from Insee (Institut National de la Statistique et des Etudes Economiques).³

Figures 1, 2 and 3 detail, in order, the spatial distribution of the total population in each French "département", the total population of immigrants, immigrants from Portugal, Morocco, Algeria, Italy, Spain, Germany, United Kingdom, and Senegal.⁴

Focusing on economic factors, two types of behavior may be expected from immigrants. Typically, they may target places with intense economic activities and, in that case, present a distribution similar to the whole population in the country; or they may form a population cluster in richer places. However, several other factors can determine the distribution of these immigrants, such as historical factors, amenities and proximity to home country. For example, Spanish immigrants settle closer to Spain, as do Italians to Italy. German people settle close to Alsace and Lorraine, following historical ties. These evaluations are clear cut in the econometric analysis.

Nonetheless, it is interesting to notice that immigrants present very uneven distribution not only when compared to the total population but also between immigrant populations themselves. On the one hand, regions with

³French institute of statistics. Source: http://www.insee.fr/fr/themes/tableau.asp?ref_id=NATnon02145®_id=0

⁴These nationalities were chosen for their historical bonds and significant presence in France.

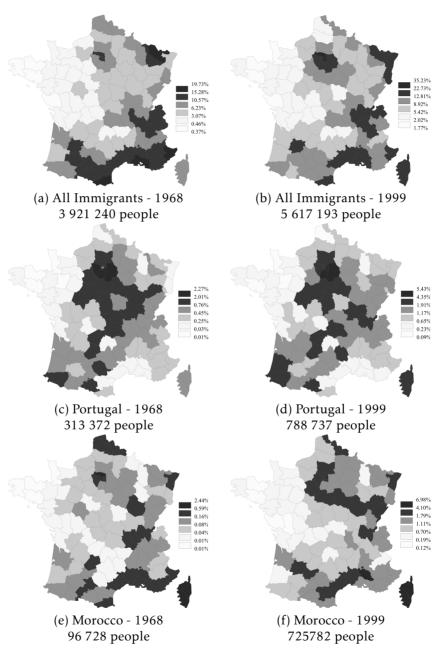


Figure 2: Geographical Distribution of Immigrants in France - All Immigrants, Portugal and Morocco

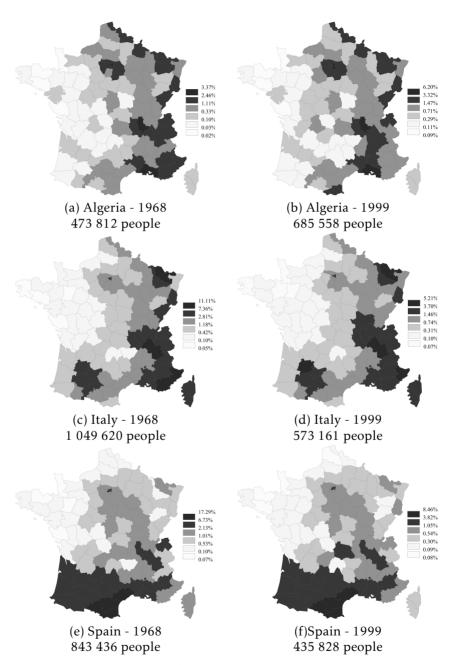


Figure 3: Geographical Distribution of Immigrants in France - Algeria, Italy and Spain

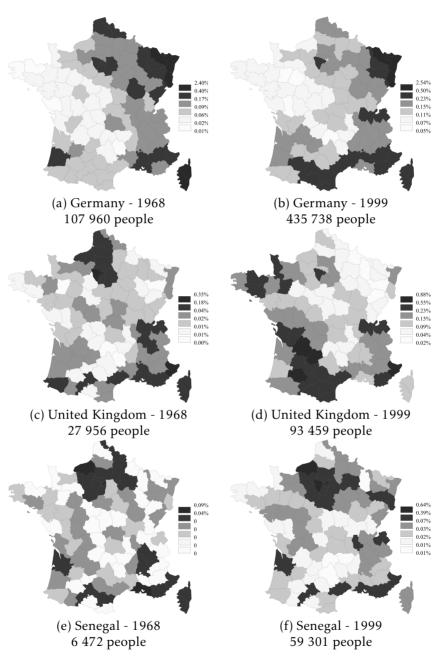


Figure 4: Geographical Distribution of Immigrants in France - Germany, United Kingdom and Senegal

intensive economic activities or rich regions seem to be preferred. On the other hand, immigrants seem to agglomerate and the first enclaves appear, over time, to determine the compatriots' location in a very robust way, even in areas that do not present economic attractiveness.

The maps in figures 1, 2 and 3 provide graphic evidence of the over-concentration phenomenon approached in the introduction and addressed by the literature (Chiswick & Miller 2004, Chiswick et al. 2002, Funkhouser 2000, Gonzalez 1998, Carrington et al. 1996). Immigrants do not pursue exclusively economic opportunities, as they also form communities.

For example immigrants from Senegal settled mainly in Paris and in the Seine-Maritime (North West of Paris) in 1968. Thirty-one years later other communities were formed, but those two agglomerations gave origin to a large network of compatriots. It is impressive how many new immigrants are settled in this region when we compare to all other distributions.

Data on French trade on regional level is available⁵ from the French Ministry of Ecology, Energy, Sustainable Development and Town and Country Planning.⁶ The data is available online for a restricted sample of countries from 2003 to 2004. Namely, these countries are Australia, Austria, Belgium, Cyprus, Czech Republic, Denmark, Estonia, Finland, Germany, Greece, Hungary, Ireland, Lithuania, Latvia, Malta, The Netherlands, Poland, Portugal, Slovakia, Spain, Sweden, Switzerland and the United Kingdom. The year 2004 is chosen for this study as it is close to the year 1999, which is the last year that migration data is available for France and it is also a more complete dataset compared to 2003.

Geographical variables such as "common border" (a dummy variable set to 1 for pairs of countries that share a border) and "common language" (dummies equal to one if both partners share a language) are extracted from the CEPII database.⁷

The data on Gross Domestic Product and national population are taken from the World Bank's "World Development Indicators".

3 Results

The potential social interaction of an individual of nationality k living in region i can be defined by:

$$S_i^k = ln\left(\sum_{j=1}^{J} \frac{s_j^k}{d_{ij}}\right) \tag{1}$$

where:

dij is the distance between region i and region j, and I assume *dij*=1,

 s_i is the share of immigrants of nationality k living in region j.

For the sake of clarity time subscripts are omitted. This section combines trade data (available on regional level) and immigration data (available on "department" level); therefore geographic location is referred to by the word "region", responding to either a French "département" or region.

⁵http://www.statistiques.equipement.gouv.fr/rubrique.php3? id_rubrique=402

⁶Ministère de l'Écologie, de l'Energie, du Développement durable et de l'Aménagement du territoire

⁷http://www.cepii.fr/anglaisgraph/bdd/distances.htm

This social component is based on the actual share of resident compatriots in French regions at a specific point in time. Assuming the inverse relationship between distance and social interaction, each region's share of immigrants in France is counted in this variable weighting the inverse of the distance between the trading region and the region where he or she lives. Immigrants living the trading region are weighted by distance equals to one. The choice to use the inverted distance follows the same functional form of many applications in economics that decreases with the distance, such as the market potential.

The estimate gravity equation of trade between country k and French region i is:

$$\ln(Trade_{kt}) = \alpha_0 + \alpha_1 S_i^k + \alpha_2 \ln(dist_{ki}) + \alpha_3 F E_k + \alpha_4 F E_i$$
(2)

where

 $dist_{ki}$ is the distance between country k and French region i

FE are fixed effects for country *k* and French region *i*

Table 1 reports results from a trade analysis. These regressions follow the last advancements in trade literature. The Poisson Pseudo Maximum Likelihood estimator is applied (Santos Silva & Tenreyro 2006, Anderson & Wincoop 2003) and specific country and region effects are controlled by fixed effects. Exports and imports are regressed separately.

Regression (1) and (2) estimates the impact of the Compatriot's Network on exports and imports, respectively. In these regressions, the distance shows an expected coefficient (Disdier & Head 2008). The immigrants' network presents a positive impact on trade. Its coefficient is significant at 1% significance level. The impact of migration on imports is higher than the impact on exports. An increase by 10% in the national stock of immigrants raises exports by 6.7% and imports by 11.2%.

Literature does not converge in this qualitative result. While Head & Ries (1998) and Wagner et al. (2002) find a more important impact of migration on imports than on exports, Gould (1994) and Girma & Yu (2002) find the opposite.⁸

Two explanations are commonly presented for higher effects of migration on imports. Firstly, immigrants may prefer consuming products from their home country (Head & Ries 1998), thus, besides the network effect (impacting both flows), this "preference" effect would only impact imports. Yet, it is assumed the network effect is both symmetric for imports and exports.⁹

Secondly, imports and exports are composed differently, according to the degree of differentiation of goods (Rauch & Trindade 2002, Tai 2009) and the effect of networks is sensitive to this variation. Yet, this bias can be inverted and also justify higher effects of migration on exports than on imports.¹⁰

A third explanation can be added to this study. A firm can somehow centralize its imports in a given county in France and then redistribute them to other counties. In this case the network variable would also capture direct

⁸See Tai (2009), pp 228-229 for a discussion of these different results.

⁹For supporting higher effects of migration on imports through preference channel, the impact of networks may also be smaller to imports than to exports.

¹⁰A definitive test for differences in the composition of imports and exports requires data on trade disaggregation on sector level, which existence cannot be proved, to the best of my knowledge.

consumption. In the three cases above, results can be upward biased for import regressions, and a measuring cannot be provided for the bias influence. Nonetheless, the effect of migration is also very significant on exports, which provides evidence of the existence of the network effect on trade.

Regressions (3) and (4) apply the instrumental variable method. They correct for a potential endogeneity between trade flows and migration. The bias is already mitigated by the use of Immigrant's stock instead of flows. However, the instrumental variable with lag of 8 years can offer a more accurate result. Coefficients are positive and significant at the 99% confidence level. Magnitudes rarely change and the coefficient for imports remains higher than the coefficient for exports. A 10% growth in the national stock of immigrants increases exports by 8.5% and imports by 16.6%.

Regressions (5) to (10) provide a comparison to other studies on the impact of migration on trade. Regressions (5) to (8) introduce the quantity of compatriots living in the reference region. It is exactly the approach typically used in this kind of literature: the total quantity of immigrants living in a country determines the bilateral trade of the country. From this estimation it is clear that this kind of procedure is missing a key part. In columns (5) and (6) the stock of migrants presents a non-significant coefficient when the Compatriot's Network is controlled. This means the network variable incorporates more information than the stock variable. Regressions (7) and (8) do not control the Compatriot's Network. Even in this case, coefficients are smaller than the network ones.

Regressions (9) and (10) apply the quantity of compatriots living in regions that share a common border. This approach is very close to that of Herander & Saavedra (2005).¹¹ It is possible to see this variable is only significant to imports at a 5% significance level. In contrast, the Compatriot's Network variable remains significant at 1% significance level for exports and imports. Yet again, the method for counting immigrants' network proposed in this study seems to give better explanation of the impact of these networks on trade.

3.1 Discussion

These results seem to confirm the strength of social interdependences and the existence of interactions even at distance. If new immigrants counted on settled compatriots' community for their installation in a host country, then networks would also have an impact on the location of foreigners. The remainder of this subsection corroborates this effect.

A first issue in the empirical analysis is the control of agglomeration forces other than social interactions. For example, Paris is a very centralizing city in France, as confirmed by many studies cited in section 2. Exogenous economic factors or some amenities could explain the massive concentration of immigrants in this region. The control featured herein considers origin country fixed effects interacting with years of fixed effect.

Discrete choices as those of immigrants for any French region are frequently estimated with a multinomial logit estimator. Instead, this work agrees with Guimarães et al. (2003), which shows that the Poisson estimator

¹¹Unlike Herander & Saavedra (2005) all out-of-region immigrants were not considered, but only border sharing immigrants, otherwise variability in that would be extremely low in cross-section data.

| Trade Regressions |
|-------------------|
| ÷ |
| Table |

| Regression Dependent (1) Exports Variable | (1) Exports | (2) Imports | (3) Exports | (4) Imports | (2) Imports (3) Exports (4) Imports (5) Exports (6) Imports (7) Exports (8) Imports (9) Exports (10) Imports | (6) Imports | (7) Exports | (8) Imports | (9) Exports | (10) Imports |
|---|---------------------------|---|-------------------------|--|--|--------------------|--|---|---|-------------------------|
| Specification | PPML | ML | | | | PPML-IV | L-IV | | | |
| Ln Compatriots' Net | 0.67^{***} (0.14) | $\begin{array}{c} 1.12^{***} \\ (0.19) \end{array}$ | 0.85^{***} (0.24) | 1.66^{***} (0.40) | 0.97^{**} (0.40) | 1.96^{***} | | | $\begin{array}{c} 0.81^{***} \\ (0.28) \end{array}$ | 1.28^{***} (0.46) |
| Ln Immig Stk within Region | | | | | -0.06 (0.13) | -0.15 (0.24) | $\begin{array}{c} 0.17^{**} \\ (0.07) \end{array}$ | $\begin{array}{c} 0.39^{***} \\ (0.13) \end{array}$ | | |
| Ln Immig Stk Neigh- boring Regions | | | | | | | | | $\begin{array}{c} 0.04 \\ (0.15) \end{array}$ | 0.68^{**} (0.29) |
| Ln Distance | $^{-0.46^{***}}_{(0.08)}$ | -0.47^{***} | -0.77^{***} | -0.49^{**} | -0.79^{***} | -0.54^{***} | -0.86^{***} | $^{-0.64^{***}}_{(0.18)}$ | -0.76^{***} | -0.129 |
| Constant | 11.46^{***} (0.76) | 14.34^{***} (1.00) | 14.78^{***} (1.39) | $ \begin{array}{c} 18.22^{***} \\ (3.13) \end{array} $ | 15.73^{***} (2.81) | 20.66*** (4.98) | 10.24^{***} (1.39) | 6.18*** (2.23) | 6.77^{***} (1.87) | 14.36^{***} (2.69) |
| Observations | 525 | 523 | 525 | 523 | 525 | 523 | 525 | 523 | 525 | 523 |
| Note: Robust standard errors in parentheses with *, ** and *** respectively denoting significance at the 1%, 5% and 10% levels. IV variables are the networks | errors in parer | ntheses with *, | ** and *** resp | ectively denot | ting significan | ce at the 1%, 5 | % and 10% le | vels. IV variat | oles are the ne | tworks |

b b 0 5 , , variables lagged once, which corresponds to 8 years averagely. provides the same coefficient and log-likelihood of multinomial logit, avoiding problems with non-linearity.

Then the estimate equation is:

$$q_{ij}^{t} = exp\left(\beta_{1}\sum_{j\neq r}^{J} ln\left(\frac{q_{ir}^{t-1}}{d_{jr}}\right) + \beta_{2}ln(distance_{ij}) + \Omega_{jt} + \Omega_{it} + Mills + cte\right)$$
(3)

where:

 q_{ij}^{t} is the quantity of immigrants from country *i* settling in French "départment" *j* in time *t*,

 $\sum_{j \neq r}^{J} ln\left(\frac{q_{ir}^{t-1}}{d_{jr}}\right)$ is the sum of settled immigrants in the previous period, weighted

by the inverse of the distance (excluding those living in region j),

 $distance_{ij}$ is the distance between country i and French "départment" j,

 Ω_{it} and Ω_{it} are fixed effects for region-time and country-time.

This section uses only immigration data (available at "departément" level), therefore geographic location is referred to by the word "département".

While country of origin *i* may be any country in the world, estinations "département" *j* are all restricted to one country, in this case France. This implies a selection bias since each immigrant who had chosen region *j* had previously chosen France as a destination country. This bias is corrected by a probit estimator. The variable Mills is the inverse of the Mills ratio which has been estimated from the first stage probit¹² estimator, following the Heckman (1979) method.

The compatriot's network is hypothesized to capture the structure that resident immigrants can offer to a new one, such as assistance with bureaucracy, language, housing, employment and business opportunities, access to home produce and leisure. This phenomenon can operate at a distance from one region to another, assuming that the new location is either deemed to be more attractive or less saturated.

Regression (1) verifies the positive impact of the Compatriot's Network on the immigrant's location. Even controlling for all specific geographic effects over time, the quantity and the proximity of pre-existing people from the same nationality determine the distribution of immigrants. This sort of agglomeration offers benefits to communication, housing, job and business possibilities, access to home produce, and leisure. These effects were analyzed by some recent studies (eg. Chiswick & Miller 2004), pointing to the evidence of a more intense concentration of immigrants, in contrast to concentration of the national population.

Regression (2) introduces country-"département" fixed effects. This method controls for all effects that are specific to each pairing of a country and "département". Geographical fixed effects are not interacted with time fixed effects in this regression because of technical limitation. The coefficient of the network measures the impact within each geographical pair considering just

¹²The selection variable is the distance between the countries of origin and France. Other independent variables of the first step probit are: the log of origin GDP, the log of origin population, dummies for common border, common language and colonial links and fixed effects for "départements" and time.

| Regression | 1 | 2 | 3 |
|--|---|-------------------------|--------------------------|
| Ln Compatriot's Network | 1.17^{***} (0.08) | 0.28^{***} (-0.03) | 0.27^{***} (-0.08) |
| Ln Distance | -0.88^{***} (-0.13) | () | -0.56^{***} (-0.07) |
| Nonselection Hazard | $ \begin{array}{c} 0.35 \\ (3.05) \end{array} $ | -0.46^{***} (0.11) | -7.35^{**} (3.43) |
| Country, "département", year F.E. | Yes | Yes | Yes |
| Country and "département" inter- acted with year F.E. | Yes | No | No |
| Country-region F.E. | No | Yes | No |
| Observations | 39805 | 30292 | 29165 |

Table 2: Choice of Location by Immigrants

Note: Robust standard errors in parentheses with *, ** and *** respectively denoting significance at the 1%, 5% and 10% levels.

the time variation of the network variable. If the quantity of resident immigrants of a given nationality, living close to a reference "département", raise by 10% each year, or if the same quantity concentrates 39.8 km each year (if each immigrant moves 39.8 km closer to the reference "département"), the quantity of immigrants living in the reference "département" raises by 2.8%.

Regression (3) corrects for endogeneity¹³ applying lagged¹⁴ variables for networks with two steps IV method. This regression includes the 16 years lagged variable, which is in fact 24 years lagged to the dependent variable. This period of time is more than enough to control endogeneity, as discussed in the appendix. The coefficient is smaller than before, though still positive and significant.

Taking column (5), a 10% increase in the network close to a reference "département" implies an increase by 2.7% in the stock of immigrants of this "département". These regressions confirm the existence of social interactions of migrants above any specific effect or endogeneity. Immigrants count on compatriots' network when deciding a location in the host country. This natural choice implies economic consequences, since the network assistance includes not only help in the moving process, but also business and job developments.

4 Conclusion

This paper investigates the extent to which social interactions impact international trade, applying data from French census distinguishing immigrants by their nationality, on the "département" level (95 "départements"), for five years.¹⁵

These interactions provide business opportunities that are verified by a very significant impact of networks on international trade. Trade of a certain region is determined not only by the social interactions of immigrants within the region, but also by the social interactions of the whole network of immigrants living in the country. This network measuring is more robust than the measuring drawn from previous research.

¹³See the appendix to details about the instrumental variable.

 $^{^{14}}$ Because data are provided by 1968, 1975, 1982, 1990 and 1990 census, one lag represents 8 years in average.

¹⁵These data are aggregated by regions (21 regions) to correspond to trade data.

A function for the role of the distance on social interactions is presented and empirical outputs show that such interactions diminish following the inverse of the distance between two immigrants. Therefore, an immigrant benefits from the compatriot's network in the destination region and also from the network installed in other regions of the country.

Results also show that the choice of location by an immigrant depends heavily on the residing network from which he or she can benefit, even when destination region, origin country, and time specific factors are controlled.

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Appendix A A Description of the Instrumental Variables for determining immigrants' location

There are two main problems concerning endogeneity. The first one is due to the use of "stocks of immigrants" as a dependent variable. As Figure 4 shows, part of this stock (square in gray) is already present at period *t*. Thus, the network variable based on the stocks of neighbor "départements" at t has a reverse causality relationship with the dependent vector. This would lead to over-estimated coefficients for the network variable. Such endogeneity is corrected by the instrumental variable method with the two lags of the network variable. This corresponds to 16 years¹⁶ and, based on the data, there is no intersection between the stock of immigrants at *t*+1 and the stock of immigrants at *t* - 2.¹⁷

The second issue is that the network variable does not take into account the stock of immigrants for the reference "départements". In Figure 4, the network variable does not sum the immigrants living in Paris at the period *t*, which could cause an over-estimation by a missing agent. However, since the reverse causality is controlled for, this effect captures the impact of immigrants living in the reference "départements" at period t who do not live there at the next period (white square below grey square). Instead, these immigrants are part of the resident network and should thus be counted. This second issue allows an improvement of the estimation as it partially considers the network within the region.

Another downside related to endogeneity derives from not considering explanatory variables that may have a significant effect. However, this is mitigated by using origin and destination fixed effects, which can be interacted with time fixed effects.

¹⁶Migration data is obtained from the French census of 1968, 1975, 1982, 1990 and 1999. Each lagged variable corresponds to 8 years difference on average.

¹⁷This conclusion is made based on data from the Institut National d'Etudes Démographiques, from France (http://www.ined.fr/): The sum of immigration flows from 1999 to 2005 is 119 422, the respective difference in the stocks is 652906. Of the total population of immigrants in 2005 (4 959 000) 31.98% were not in France in 1999. On average, each year from 1999 to 2005 6% of the population is composed of new immigrants (immigrants who arrived in the current year).

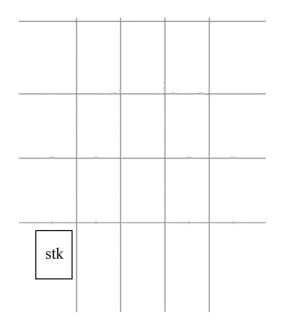


Figure A.1: Endogeneity and the Stock of Immigrants